

Uniform distⁿ

X taking values x_1, \dots, x_n .

$$P(X = x_i) = \frac{1}{n} \quad \forall i$$

Poisson Distⁿ.

$$f(x) = e^{-\lambda} \frac{\lambda^x}{x!}, \quad x=0, 1, 2, \dots$$

$$E(X) = \sum_{x=0}^{\infty} x e^{-\lambda} \frac{\lambda^x}{x!}$$

$$\begin{aligned} &= \lambda \sum_{x=1}^{\infty} e^{-\lambda} \frac{\lambda^{x-1}}{(x-1)!} \\ &= \lambda \sum_{y=0}^{\infty} e^{-\lambda} \frac{\lambda^y}{y!} = \lambda \end{aligned}$$

$$\begin{aligned}
 \text{Var}(x) &= \sum_{x=1}^{\infty} x^2 \frac{e^{-\lambda} \lambda^x}{x!} - \lambda^2 \\
 &= \cancel{x} + \cancel{-\lambda} \\
 &= \cancel{\lambda} \\
 &= \sum_{x=1}^{\infty} x \sum_{y=1}^{\infty} y \cdot \frac{e^{-\lambda} \lambda^x}{(x-y)!} - \lambda^2 \\
 &= \sum_{y=0}^{\infty} \underbrace{y e^{-\lambda} \lambda^y}_{y=0} \sum_{x=y+1}^{\infty} \frac{(y+1)}{(x-y)!} \frac{e^{-\lambda} \lambda^x}{x!} - \lambda^2 \\
 &= \underbrace{\sum_{y=0}^{\infty} y e^{-\lambda} \lambda^y}_{y=0} \underbrace{\sum_{x=y+1}^{\infty} \frac{(y+1)}{(x-y)!} \frac{e^{-\lambda} \lambda^x}{x!}}_{y!} - \lambda^2
 \end{aligned}$$

Defⁿ Given two d.v. X and Y ,
the Covariance of X and Y , denoted
by $\text{Cov}(x, y)$ is given by

$$\text{Cov}(x, y) = E(XY) - E(X) \cdot E(Y)$$

Thm Let x, r be $\sigma \cdot v + c, b$

constants. Then

$$\text{Var}(ax+by) = a^2 \text{Var}(x) + b^2 \text{Var}(y) + 2ab \text{Cov}(x, r).$$

Pf. Define $E(x) = \mu_x$ & $E(r) = \mu_y$.

$$E(ax+by) = a\mu_x + b\mu_y.$$

$$\begin{aligned}
 \text{Var}(ax + by) &= \{ (ax + by)^2 - (a\mu_x + b\mu_y)^2 \} \\
 &= E(a^2x^2 + b^2y^2 + 2abxy) - (\tilde{a}\mu_x^2 + \tilde{b}\mu_y^2 + 2ab\mu_x\mu_y) \\
 &= a^2(E(x^2) - \mu_x^2) + b^2(E(y^2) - \mu_y^2) \\
 &\quad + 2ab(E(xy) - E(x)E(y)) \\
 &= a^2\text{Var}(x) + b^2\text{Var}(y) + 2ab\text{Cov}(x, y).
 \end{aligned}$$

Ex. Generalise this to more than 2 r.v.s.

Ex. Let X be a binomial d. v. with parameters n and p .

$$X_i = \begin{cases} 1 & \text{with prob. } p \\ 0 & \text{with prob. } q = 1 - p. \end{cases}$$

$$X = \sum_{i=1}^n X_i, \quad \bar{t}(X) = \sum_{i=1}^n \bar{t}(X_i) = np.$$

$$V_{\text{var}}(x_i) = 1^2 p + 0 \cdot q - p^2 \\ = p(1-p) = pq.$$

$$V_{\text{var}}(x) = \sum_{i=1}^n V_{\text{var}}(x_i) = npq.$$

Suppose there are n points
 $(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)$ in
the plane. Given a positive integer m ,
which polyⁿ of degree m is "closest"
to these points. At the desired

poly

$$p(x) = a + \sum_{i=1}^m b_i x^i$$

The polyⁿ What is the "best"
is the polyⁿ whose coeff's
minimizes the following

$$L = \sum_{i=1}^n (y_i - p(x_i))^2$$

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Thm 1. Given n points $(x_1, y_1), \dots, (x_n, y_n)$.

The st. line that minimize L .

has gradient

and y -intercept

$$b = \frac{\sum_{i=1}^n x_i y_i - n \bar{x} \bar{y}}{\sum_{i=1}^n x_i^2 - n (\bar{x})^2}$$

$$a = \bar{y} - b \bar{x},$$

where $\bar{x} = \frac{1}{n} \sum x_i$

$$\bar{y} = \frac{1}{n} \sum y_i$$

Pf

Since (a, b) minimizes L , we

must have $\frac{\partial L}{\partial b} = 0$

& $\frac{\partial L}{\partial a} = 0$.

We have

$$L = \sum_{i=1}^n (y_i - (a + bx_i))^2$$

$$\frac{\partial L}{\partial b} = \sum_{i=1}^n 2(y_i - a - bx_i)(-x_i)$$

$$\frac{\partial L}{\partial b} = 0 \quad ; \text{implies}$$

$$a \sum_{i=1}^n x_i + b \sum_{i=1}^n \bar{x}_i = \sum_{i=1}^n x_i y_i$$

①

$$\begin{aligned} n a + b n \bar{x} &= n \bar{y} \\ a + b \bar{x} &= \bar{y} \Rightarrow \boxed{a = \bar{y} - b \bar{x}} \end{aligned}$$

from ①

$$\begin{aligned} (\bar{y} - b \bar{x}) n \bar{x} + b \sum_{i=1}^n x_i^2 &= \sum_{i=1}^n x_i y_i \\ b \left[\sum_{i=1}^n x_i^2 - n \bar{x}^2 \right] &= \sum_{i=1}^n x_i y_i - n \bar{x} \bar{y} \\ b = & \frac{\sum_{i=1}^n x_i y_i - n \bar{x} \bar{y}}{\sum_{i=1}^n x_i^2 - n \bar{x}^2} \end{aligned}$$

Exponential Regression

Suppose x and y are related by an exponential fit of the form.

$$y = a e^{bx}$$

$$\log y = \log a + bx. \quad \text{between } x \text{ & } \log y$$

Hence the linear relation that is the best fit is given by

$$\log y = \log a + bx \rightarrow \text{where}$$

$$b = \frac{\sum_{i=1}^n x_i \log y_i - \bar{x} \cdot \sum_{i=1}^n \log y_i}{\sum_{i=1}^n x_i^2 - n \bar{x}^2}$$

$$\log a = \frac{1}{n} \sum_{i=1}^n \log y_i - b \bar{x}$$

Log regression.

$$y = a x^b$$

$$\log y = \log a + b \log x.$$



\downarrow
 x .

Logistic regression.

$$y = \frac{L}{1 + e^{a+bx}}, \quad L, a, b \text{ are const.}$$

$$1 + e^{a+bx} = \frac{L}{y}.$$

$$e^{a+bx} = \frac{L-y}{y}.$$

$$\left\{ \begin{array}{l} a+bx = \log \left(\frac{L-y}{y} \right) \end{array} \right.$$



Correlation coefficient

Let x, y be $\sigma.$ v. with means μ_x, μ_y
and S.Ds. σ_x, σ_y respectively

Put $X^* = \frac{x - \mu_x}{\sigma_x}$

$$Y^* = \frac{y - \mu_y}{\sigma_y}$$

We define the correlation coeff by

$$\rho = \text{Cov}(X^*, Y^*) = E\left(\frac{X - \mu_X}{\sigma_X} \cdot \frac{Y - \mu_Y}{\sigma_Y}\right) - E\left(\frac{X - \mu_X}{\sigma_X}\right) E\left(\frac{Y - \mu_Y}{\sigma_Y}\right)$$

Proof. $|\rho| \leq 1$.

Pf

$$O \leq E((x^* + \gamma^*)^2)$$

$$= E(x^{*2}) + E(\gamma^{*2}) + 2E(x^*\gamma^*).$$

$$= E(x^{*2}) + E(\gamma^{*2}) + 2S + 2E(x^*)E(\gamma^*).$$

$$x \in \mathbb{R}^n$$

$$y \in \mathbb{R}$$

$$x_i = x_i - \mu_x$$

$$0 \leq \left(\sum_{i=1}^n x_i y_i \right)^2 \leq \left(\sum_{i=1}^n x_i^2 \right) \cdot \left(\sum_{i=1}^n y_i^2 \right)$$

$$x_i - \mu_x$$